

Sub - SPT Set (A) CSE (ivsem)

Q-1 A and B take turn in throwing two die on the understanding that the first two throw 9 will be awarded prize. If A has first turn show that their chance of winning are in the ratio 9:8

Sol - Two die may have $6 \times 6 = 36$ ways.

$$P = \frac{4}{36} = \frac{1}{9} \quad q = 1 - P = 1 - \frac{1}{9} = \frac{8}{9}$$

Prob of 9 at first throw = P

... at second = qP

... Third = q^2P

Hence A's chance of winning = chance of 9 occurring all the first at same 1, 3, 5th. . .

$$\begin{aligned} P(A) &= P + q^2P + q^4P + \dots \\ &= \frac{P}{1-q^2} \rightarrow (1) \end{aligned}$$

Similarly

$$P(B) = qP + q^3P + q^5P + \dots = \frac{qP}{1-q^2}$$

$$\frac{P(A)}{P(B)} = \frac{P}{qP} = \frac{1}{q} = \frac{9}{8}$$

$$\boxed{P(A) : P(B) = 9:8} \quad \text{Ans}$$

or

Q-2 Q-1 A Continuous random variable X that can be assume any value between $x=1$ and $x=4$ and is zero elsewhere has the density function given by $f(x) = \lambda(1+x)$. Find $P(X < 3) \geq f$

Sol - 1

we know that $\int_{-\infty}^{\infty} f(x) dx = 1$

$$= \int_{-\infty}^{-1} f(x) dx + \int_{-1}^4 f(x) dx + \int_4^{\infty} f(x) dx$$

$$= 0 + \int_{-1}^4 f(x) dx + 0 = \int_{-1}^4 \lambda(1+x) dx$$

$$= [\lambda x + \frac{\lambda x^2}{2}]_{-1}^4 = 1$$

$$\lambda [3 + (\lambda - \frac{1}{2})] = 1 \Rightarrow (10 + \lambda) \lambda = 1$$

$$\lambda = \frac{2}{21}$$

$$F(3) = P(X \leq 3) = \int_{-\infty}^3 f(x) dx$$

$$= \frac{2}{21} \int_{-1}^3 (1+x) dx = \frac{2}{21} \left[x + \frac{x^2}{2} \right]_{-1}^3$$

$$= \frac{2}{21} \left[2 + (\lambda - \frac{1}{2}) \right] = \frac{12}{21} = \frac{4}{7} \underline{\underline{\text{Ans}}}$$

Q-2 A random variable X has the prob distribution

$$\begin{array}{cccccccc} X & = & 0 & 1 & 2 & 3 & 4 & 5 & 6 & 7 \\ P(X) & = & 0 & K & 2K & 2K & 3K & K^2 & 2K^2 & 7K^2 + K \end{array}$$

Find K and $P(1.5 < X < 4.5)$

$X \geq 2$

Sol we have $\sum_{x=0}^7 P(x) = 1$

$$K + 2K + 2K + 3K + K^2 + 2K^2 + 7K^2 + K = 1$$

$$10K^2 + 9K - 1 = 0$$

$$K = -1 \quad (\text{neglect}) \quad K = -1, K = \frac{1}{10}$$

$$\text{take } K = \frac{1}{10}$$

$$\begin{aligned}
 P\left(\frac{1.5 < X < 4.5}{X \geq 2}\right) &= \frac{P(2 < X < 4.5)}{P(X \geq 2)} \\
 &= \frac{P(X=3) + P(X=4)}{P(X \geq 2)} \\
 &= \frac{2 \cdot \frac{1}{10} + 3 \cdot \frac{1}{10}}{1 - [P(X=0) + P(X=1) + P(X=2)]} \\
 &= \frac{\frac{5}{10}}{1 - \left(\frac{3}{10}\right)} = \underline{\underline{5/7}}
 \end{aligned}$$

Q2 A r.v. X has the following density function

$x =$	-2	-1	0	1	2	3
$P(X=x) =$	0.1	K	0.2	$2K$	0.3	$3K$

Sol. X is a random variate

$$\sum_{i=-2}^3 P_i = 1$$

$$0.1 + 0.2 + 0.3 + 6K = 1$$

$$6K = 0.4$$

$$K = \frac{1}{15}$$

$$\begin{aligned}
 E(X) &= \sum p_i x_i = (-2)(0.1) - K + 0 + 2K + 0.6 + 9K \\
 &= 10K + 0.4 \\
 &= \frac{10}{15} + \frac{4}{10} \\
 &= \underline{\underline{\frac{16}{15}}}
 \end{aligned}$$

Q8-3 Express the probability of Poisson distribution
its M.g.f moment about origin, central moment
variance & S.D

So Poisson distribution apply the following rules

(i) $p \rightarrow 0$ & $n \rightarrow$ very large.

$$\text{prob of } r \text{ successes} = \lim_{\substack{p \rightarrow 0 \\ np = m}} n_{cr} p^r (1-p)^{n-r}$$

$$P(X=r) = \lim_{\substack{n \rightarrow \infty \\ np = m}} \frac{\{n(r-1) \dots (n-r+1)\}^p}{L^r} (1-\frac{m}{n})^n (1-\frac{m}{n})^r$$

$$\text{Let } t = (1-\frac{m}{n})^n$$

$$\log(1-n) = -n - n^2/2 - n^3/3 - n^4/4 - \dots$$

$$\log t = m \log(1-\frac{m}{n}) = m \left[-\frac{m}{n} - \frac{m^2}{2n^2} - \frac{m^3}{3n^3} - \dots \right]$$

$$= -m - \frac{m^2}{2n} - \frac{m^3}{3n^2} - \dots$$

$$t = e^{-m - \frac{m^2}{2n} - \frac{m^3}{3n^2}} \dots$$

$$= e^{-m}$$

$$P(X=r) = \left[\frac{(np)^r}{L^r} (1-\frac{1}{n})(1-\frac{2}{n}) \dots (1-\frac{r-1}{n}) \right] e^{-m}$$

$$\boxed{P(X=r) = \frac{m^r}{L^r} e^{-m}} \quad r=0, 1, 2, 3, \dots$$

M.g.f of Poisson distribution:

$$\begin{aligned} M_x(t) &= E(e^{tx}) \\ &= \sum_{r=0}^{\infty} e^{-m} \frac{m^r}{L^r} e^{tr} = \sum_{r=0}^{\infty} e^{-m} \frac{(me^t)^r}{L^r} \\ &= e^{-m} (e^{met}) \end{aligned}$$

$$\text{we know } \bar{x} = \mu = \left[\frac{d}{dt} M_x(t) \right]_{t=0} = \left[e^{-m} (e^{met}) me^t \right]_{t=0} = m$$

$$M_2' = \left[\frac{d^2 M_0(t)}{dt^2} \right]_{t=0} = e^m \left[(e^{mt})^2 + e^{2mt} \right]_{t=0}$$

$$= e^{-m} e^m [m^2 + m] = m^2 + m$$

\therefore Standard deviation $\sigma^2 = M_2' - M_1'^2 = m^2 + m - m^2 = m$

$$\boxed{S.D. \text{ is } \sigma = \sqrt{m}}$$

B-3 Define binomial distribution, m.g.f., mean, variance
Q S.D. ?

Sol Probability p_i of r success and $(n-r)$ failure
in n trials is

$$p_i = P(X=r) = {}^n C_r q^{n-r} p^r$$

$$= \sum_{r=0}^n {}^n C_r p^r q^{n-r}$$

$$= n_r q^n + n_{r-1} q^{n-1} p + \dots + n_{C_n} p^n$$

It is binomial expansion of $(q+p)^n$

$$\sum_{r=0}^n P(X=r) = \sum_{r=0}^n p_i = (q+p)^n = 1$$

M.g.f. of binomial distribution -

M.g.f. of binomial distribution about origin
= expected value of (e^{tr})

$$M_r(t) = E(e^{tr})$$

$$M_r(t) = \sum_{r=0}^n p_r e^{tr}$$

$$= \sum_{r=0}^n {}^n C_r q^{n-r} p^r e^{tr}$$

$$= \sum_{r=0}^n {}^n C_r q^{n-r} (pe^t)^r$$

$$= (q + pe^t)^n$$

$$\text{Mean } \mu_1' = \frac{d}{dt} [M_0(t)] \\ = [n(q+pe^t)^{n-1} pe^t]_{t=0}$$

$$\text{Mean} = \mu_1' = np$$

$$\mu_2' = \left[\frac{d^2}{dt^2} M_0(t) \right]_{t=0} = [n(n-1)(q+pe^t)^{n-2}(pe^t)^2 + n(q+pe^t)^{n-1}p \cdot e^t]_{t=0} \\ \mu_2' = n(n-1)p^2 + np$$

$$\sigma^2 = \mu_2 - \mu_1'^2 \\ = n(n-1)p^2 + np - (np)^2 \\ = -np^2 + np = np\{1-p\} \\ \boxed{\sigma^2 = npq}$$

S.D

$$\sigma = \sqrt{npq}$$

B-4 The sum of mean and variance of a binomial distribution is 15 and the sum of their squares is 117 determine the distribution.

Sol we have mean = np
variance = npq

$$np + npq = 15 \Rightarrow n^2 p^2 + n^2 p^2 q^2 = 117$$

$$n^2 p^2 [1+q^2] = 117$$

$$\frac{1+q^2}{1+q^2} = \frac{225}{117} \Rightarrow 1 + \frac{2q}{1+q^2} = \frac{225}{117}$$

$$\frac{2q}{1+q^2} = \frac{12}{13}$$

$$\frac{q}{1+q^2} = \frac{6}{13}$$

$$6q^2 - 13q + 6 = 0$$

$$q = \frac{13 \pm \sqrt{169 - 144}}{12} = \frac{13 \pm 5}{12} \text{ or } \frac{13 \mp 5}{12}$$

$$q = \frac{2}{3}$$

or

$$np[1+q] = 15$$

$$p = 1 - q = \frac{1}{3}$$

$$n = 27$$

Hence the required dis

$$P(X=x) = {}^{27}C_x \left(\frac{2}{3}\right)^x \left(\frac{1}{3}\right)^{27-x}$$

$$x = 0, 1, 2, \dots, 27$$

Q-4 Prob that a man aged 60 would be alive till the 70 years of age in 0.65. Find the probability that at least 7 out of 10 such men would be alive till 70 years of age

$$\underline{\text{Sol}} \quad n = 10 \quad p = 0.65 \quad q = 0.35$$

$$P(X \geq 7) = P(7) + P(8) + P(9) + P(10)$$

$$= {}^{10}C_7 (0.35)^3 (0.65)^7 + {}^{10}C_8 (0.35)^2 (0.65)^8$$

$$+ {}^{10}C_9 (0.35)^1 (0.65)^9 + (0.65)^{10}$$

$$= 120 \times 0.0210 + 45 \times 0.0390 + 10 \times 0.0725$$

$$+ 0.01346$$

$$= .252 + .1755 + .0725 + .01346$$

$$= 0.513 \quad \underline{\text{Ans}}$$

(B)

Bay's theorem:

Sol Let $\{A_1, A_2, \dots, A_n\}$ be an event space. Suppose an event A occurred but it is not known which of the events A_1, A_2, \dots, A_n holds true. Then the conditional probability that event A_k occurred given that A has been occurred is given by

$$P\left(\frac{A_k}{A}\right) = \frac{P(A_k) \cdot P\left(\frac{A}{A_k}\right)}{\sum_{j=1}^n P(A_j) \cdot P\left(\frac{A}{A_j}\right)} \quad k=1, 2, \dots, n$$

Proof: Let $\{A_1, A_2, \dots, A_n\}$ be an event space such that A_1, A_2, \dots, A_n are mutually exclusive and exhaustive events.

$$\begin{aligned} P(A_k \cdot A) &= P(A) \cdot P\left(\frac{A_k}{A}\right) \\ &= P(A_k) \cdot P\left(\frac{A}{A_k}\right) \end{aligned}$$

$$\Rightarrow P\left(\frac{A_k}{A}\right) = \underbrace{P(A_k) \cdot P\left(\frac{A}{A_k}\right)}_{P(A)}$$

But by theorem of total probability.

$$P(A) = \sum_{j=1}^n P(A_j) \cdot P\left(\frac{A}{A_j}\right)$$

$$\begin{aligned} P\left(\frac{A_k}{A}\right) &= P(A_k) \cdot P\left(\frac{A}{A_k}\right) \\ &= \frac{\sum_{j=1}^n P(A_j) \cdot P\left(\frac{A}{A_j}\right)}{\sum_{j=1}^n P(A_j) \cdot P\left(\frac{A}{A_j}\right)} \quad \underline{\text{Proved}} \end{aligned}$$

$k=1, 2, \dots, n$

Q1 The first four moment in a certain probability distribution about the point 4 are -1.5, 17, -30 and 108, calculate β_1 and β_2 and state whether the distribution is leptokurtic or platykurtic.

Sol Given

$$(M'_1)_4 = \sum p_i (x_i - 4) = -1.5$$

$$\sum p_i x_i - 4 = -1.5$$

$$\bar{x} = 4 - 1.5 = 2.5$$

$$(M'_2)_4 = \sum p_i (x_i - 4)^2 = 17$$

$$17 = \sum p_i [(x_i - \bar{x}) + (\bar{x} - 4)]^2$$

$$= M_2 + 0 + (2.5 - 4)^2$$

$$M_2 = 17 - 2.25 = 14.75$$

$$(M'_3)_4 = \sum p_i (x_i - 4)^3 = -30$$

$$-30 = \sum p_i [(x_i - \bar{x}) + (\bar{x} - 4)]^3$$

$$= M_3 + 3(1.5)(14.75) + 0 - (1.5)^3$$

$$M_3 = -30 + 9/2 \times 59/4 + 3 \cdot 375 = 39.75$$

$$(M'_4)_4 = \sum p_i (x_i - 4)^4 = 108$$

$$M_4 = 108 + 6(39.75) - 13.5(14.75) - (1.5)^4 \\ = 142.3$$

$$\beta_1 = \frac{M_3^2}{M_2^3} = \frac{(39.75)^2}{(14.75)^3} = 0.492$$

$$\beta_2 = \frac{M_4}{M_2^2} = \frac{142.3}{(14.75)^2} = 0.68 \leftarrow$$

so the curve is platykurtic.

Q-2 If the life time of a component has PDF $\lambda e^{-\lambda t}$, $\lambda > 0$, $t \geq 0$ compute its time to failure, variance and failure rate function.

Sol Time to failure is its mean life time

$$\begin{aligned} E(x) &= \int_0^\infty t f(t) dt = \int_0^\infty \lambda t e^{-\lambda t} dt \\ &= \lambda \left[\left(-\frac{t e^{-\lambda t}}{\lambda} \right)_0^\infty + \frac{1}{\lambda} \int_0^\infty e^{-\lambda t} dt \right] \\ &= \lambda [0 - \gamma_{\lambda}^2 (e^{-\lambda t})_0^\infty] = \gamma_{\lambda} \end{aligned}$$

$$\begin{aligned} M_1 &= E(x^2) = \int_0^\infty t^2 f(t) dt = \lambda \int_0^\infty t^2 e^{-\lambda t} dt \\ &= \lambda \left[\left(-\frac{t^2 e^{-\lambda t}}{\lambda} \right)_0^\infty + \frac{2}{\lambda} \int_0^\infty t e^{-\lambda t} dt \right] \\ M_2 &= \gamma_{\lambda}^2 \end{aligned}$$

Variance $\sigma^2 = M_2 - M_1^2 = 2\gamma_{\lambda}^2 - (\gamma_{\lambda})^2 = \gamma_{\lambda}^2$

Failure rate function $\lambda(t) = \frac{f(t)}{R(t)}$

$$\begin{aligned} f(t) &= \lambda e^{-\lambda t} \quad R(t) = 1 - F(t) \\ &= 1 - \int_0^t f(u) du \\ &= 1 - \lambda \int_0^t e^{-\lambda u} du \\ &= 1 - \left[\frac{1 - e^{-\lambda t}}{\lambda} \right] = e^{-\lambda t} \end{aligned}$$

$$\lambda(t) = \frac{\lambda e^{-\lambda t}}{e^{-\lambda t}} = \lambda$$

Q-2 Given the joint probability density

$$f(x,y) = \begin{cases} 2/3(x+2y) & 0 < x < 1, 0 < y < 1 \\ 0 & \text{elsewhere} \end{cases}$$

Find (i) Marginal density of X and Y

(2) Conditional density of X given $y=y$

and use it evaluate $P[X \leq 1/2 | y=y_2]$

Sol Marginal density of X is

$$\begin{aligned} g(x) &= \int_{-\infty}^{\infty} f(x,y) dy \\ &= \int_0^1 2/3(x+2y) dy = \\ &= \begin{cases} 2/3(x+1) & \text{for } 0 < x < 1 \\ 0 & \text{elsewhere} \end{cases} \end{aligned}$$

Similarly marginal density of Y is

$$\begin{aligned} h(y) &= \int_0^1 2/3(x+2y) dx \\ h(y) &= \begin{cases} 1/3(1+4y) & 0 < y < 1 \\ 0 & \text{elsewhere} \end{cases} \end{aligned}$$

(iii) By conditional density.

$$\phi(x|y) = \frac{f(x,y)}{h(y)}$$

$$\phi(x|y) = \frac{2/3(x+2y)}{1/3(1+4y)} = \frac{2(x+2y)}{1+4y}$$

So conditional density of X given by $y=y$ is

$$\phi(x|y) = \begin{cases} \frac{2(x+2y)}{1+4y} & 0 < x < 1 \\ 0 & \text{elsewhere} \end{cases}$$

$$P\left[\frac{x \leq y_2}{y=y_2}\right] = \int_0^{y_2} \frac{2(x+1)}{1+2} dx$$

$$= 2/3 \left[\frac{x^2}{2} + x \right] = 2/3 [y_2^2 + y_2] = 5/12 \text{ Ans}$$

Q-3 Show that for the exponential distribution given by $dP = \alpha e^{-\alpha x} dx$, α being constant the mean and S.D. are equal to c .

Sol: we have by P.d.f $\int_0^\infty dP = 1$

$$\int_0^\infty \alpha e^{-\alpha x} dx = \alpha (-c) [e^{-\alpha x/c}]_0^\infty \\ \Rightarrow \alpha c = 1 \quad [\alpha = \frac{1}{c}],$$

by def of moment

$$\mu_r = E(x^r) = \int_0^\infty x^r dP \\ = \int_0^\infty x^r \frac{1}{c} e^{-x/c} dx \\ = c^r \int_0^\infty u^r e^{-u} du \quad \text{if } x=cu \\ = c^r \int_0^\infty u^{(r+1)-1} e^{-u} du \\ = c^r \Gamma(r+1) = c^r L_n \quad (\text{by Gamma fun})$$

$$\text{If } r=1 \quad \mu_1 = c^1 L_1 = c$$

$$\mu_2 = c^2 L_2 = 2c^2$$

$$\sigma^2 = 2c^2 - c^2 = c^2$$

$$S.D. = \sqrt{c^2} = c \quad \text{Proven}$$

Q-3 Find P.M.F. of a random variate x whose prob generating function.

at $x=3$

$$G_x(z) = \frac{2+z}{(2-z^2)(4-z)}$$

Sol. The P.G.F is given

$$G_x(z) = (2+z) \times \frac{1}{2} \left(1 - \frac{z^2}{2}\right)^{-1} \times \frac{1}{4} \left(1 - \frac{z^4}{4}\right)^{-1}$$

$$G_x(z) = \frac{1}{8} (2+z) \left[1 + \frac{z^2}{2} + \frac{z^4}{4} + \dots\right] \left[1 + \frac{z^4}{4} + \frac{z^8}{16} + \frac{z^{12}}{64} + \dots\right]$$

$$= \frac{1}{8} [(2+z)(1 + \frac{z^2}{2} + \frac{z^4}{4} + \dots)] \left[1 + \frac{z^4}{4} + \frac{z^8}{16} + \frac{z^{12}}{64} + \dots\right]$$

$$C_X(z) = \frac{1}{8} \left[(2+z+z^2+z^3)_3 + z^2 (z^2+z^3)_4 + \dots \right] \\ \left[1+z_4 + z^2 (z^2+z^3)_6 + z^3 (z^3+z^4)_8 + \dots \right]$$

Prob at $x=3$

$$P(X=3) = \frac{1}{4 \times 64} + \frac{1}{8 \times 16} + \frac{1}{8 \times 4} + \frac{1}{8 \times 2} \\ = \frac{1}{16} \left[\frac{1}{16} + \frac{1}{8} + \frac{1}{2} + 1 \right] = \frac{27}{256} \approx$$

Q-4 Find the M.g.f of the random variable whose p.m.f is given by $P(X=x) = \frac{1}{8} C_x^3$ $x=0, 1, 2, 3$,
and Find M'_1 M'_2 ?

Sol we have $m_x(t) = E(e^{tx})$

$$= \sum_{x=0}^3 e^{tx} \cdot \frac{1}{8} C_x^3 \\ = \frac{1}{8} [1 + e^t + 3e^{2t} + e^{3t}] \\ = \frac{1}{8} [1 + e^t]^3$$

$$M'_1 = \left[\frac{d}{dt} m_x(t) \right] = \frac{3}{8} [(1+e^t)^2 e^t]_{t=0} \\ = \frac{3}{2}$$

$$M'_2 = \frac{d^2}{dt^2} [m_x(t)] = \frac{3}{8} \left[(1+e^t)^2 e^t + 3 \frac{d}{dt} (1+e^t)^2 e^t \right]_{t=0} \\ = \frac{3}{2} + \frac{3}{2} = 3$$

B-4 Define exponential distribution, M.g.f, mean, Variance & S.D

Def A continuous variate X such that $0 \leq X < \infty$ is said to follow the exponential distribution when its prob density function is given by

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

we have $\int_0^\infty f(x) dx = \lambda \frac{[e^{-\lambda x}]_0^\infty}{-\lambda} = 1$

M.g.f:

$$\begin{aligned} M_X(t) &= E[e^{tx}] = \int_0^\infty f(x) e^{tx} dx \\ &= \lambda \int_0^\infty e^{-(\lambda-t)x} dx \\ &= \lambda \left[\frac{e^{-(\lambda-t)x}}{-(\lambda-t)} \right]_0^\infty = \frac{\lambda(0-1)}{-(\lambda-t)} = \frac{\lambda}{\lambda-t} \\ &= [1 - t/\lambda]^{-1} = 1 + 1/\lambda t + \frac{1}{\lambda^2} t^2 + \dots \end{aligned}$$

Mean $\bar{x} = \mu_1' = \left[\frac{d}{dt} [M_X(t)] \right]_{t=0} = \left[\frac{1}{(\lambda-t)^2} \right]_{t=0}$

$$= \lambda / \lambda^2 = \gamma_\lambda$$

Variance $\sigma^2 = \mu_2' - (\mu_1')^2$

$$= \left[\frac{2\lambda}{(\lambda-t)^3} \right]_{t=0} - \gamma_\lambda^2$$

$$\sigma^2 = \gamma_\lambda^2 - \gamma_{\lambda^2} = \gamma_{\lambda^2}$$

S.D $= \sigma = \sqrt{\gamma_{\lambda^2}} = \underline{\gamma_\lambda}$

**RIET****RAJASTHAN INSTITUTE OF
ENGINEERING & TECHNOLOGY**

Approved by AICTE & Affiliated to Rajasthan Technical University

I Mid Term examination

Session: 2017-18

B.Tech II Year CSE (IV Semester)

Subject with code: SPT

SET-B

Time: 2hrs.

M.M.:20

Attempt all questions

Q-1 Prove that Baye's theorem !

OR

The first four moment in a certain probability distribution about the point 4 are -1.5, 17, -30 and 108. Calculate β_1, β_2 , and state whether the distribution is lepto kurtic or platykurtic.

Q-2 If the life time of a component has probability density function $\lambda e^{-\lambda t}$. $\lambda > 0, t > 0$. compute its time to failure, variance and failure rate function.

OR

Given the joint probability density .

$$f(x,y) = \begin{cases} \frac{2}{3}(x+2y) & 0 < x < 1, 0 < y < 1 \\ 0 & \text{Elsewhere} \end{cases}$$

Find (1) marginal density of X and Y.

(2) Conditional density of X given $Y=y$. and use it to evaluate $P(X \leq 1/2 | Y = 1/2)$.

Q-3 Show that for the exponential distribution given by $dp = ae^{-x/c} dx$, and x lies between 0 to infinite , a being a constant, the mean standard deviation are equal to c ..

OR

Find the probability mass function of a random variate x whose probability generating function is

$$G_x(z) = \frac{2+z}{(2-z^2)(4-z)} \quad \text{at } X=3 .$$

Q-4 Find the m.g.f of the random variable X whose p.m.f is given by

 $P(X=x) = 1/8 C_x^3 ; x=0, 1, 2, 3$ and then find μ_1, μ_2 !

OR

Define exponential distribution and its MGF, mean, variance, SD

